

Solution #1 (2.16)

$$\frac{\partial \ln p(\mathbf{x}; \theta)}{\partial \theta} = \frac{\partial p(\mathbf{x}; \theta)}{\partial \theta} \frac{1}{p(\mathbf{x}; \theta)}$$

By the definition of the mean

$$\begin{aligned} \mu &= \int_{-\infty}^{\infty} \frac{\partial p(\mathbf{x}; \theta)}{\partial \theta} \frac{p(\mathbf{x}; \theta)}{p(\mathbf{x}; \theta)} d\mathbf{x} = \int_{-\infty}^{\infty} \frac{\partial p(\mathbf{x}; \theta)}{\partial \theta} d\mathbf{x} \\ &\stackrel{(a)}{=} \frac{\partial}{\partial \theta} \int_{-\infty}^{\infty} p(\mathbf{x}; \theta) d\mathbf{x} = \frac{\partial 1}{\partial \theta} = 0 \end{aligned}$$

Interchanging order of differentiation and integration in (a) is allowed since the limits do not depend on θ .

Solution #2

2.17 The likelihood function is

$$P(X; q) = \prod_{i=1}^N q^{x_i} (1-q)^{(1-x_i)}$$

or

$$P(X; q) = q^{\sum_{i=1}^N x_i} (1-q)^{(N - \sum_{i=1}^N x_i)}$$

$$\begin{aligned} \frac{\partial P(X; q)}{\partial q} &= \left(\sum_{i=1}^N x_i \right) q^{\left(\sum_{i=1}^N x_i - 1 \right)} (1-q)^{(N - \sum_{i=1}^N x_i)} \\ &\quad - \left(N - \sum_{i=1}^N x_i \right) (1-q)^{(N - \sum_{i=1}^N x_i - 1)} q^{\sum_{i=1}^N x_i} = 0 \end{aligned}$$

or

$$q^{\sum_{i=1}^N x_i} (1-q)^{(N - \sum_{i=1}^N x_i)} \left(\frac{\sum_{i=1}^N x_i}{q} - \frac{N - \sum_{i=1}^N x_i}{1-q} \right) = 0$$

The solutions $q = 0, 1$ result in a minimum of $P(X; q)$. The maximum comes from

$$\frac{\sum_{i=1}^N x_i}{q} - \frac{N - \sum_{i=1}^N x_i}{1-q} = 0 \Rightarrow q = \frac{1}{N} \sum_{i=1}^N x_i$$

Solution #3

2.20 The likelihood function is

$$\begin{aligned} L(\theta) &= \sum_{k=1}^N \ln \theta^2 + \sum_{k=1}^N \ln x_k - \sum_{k=1}^N \theta x_k \\ &= N \ln \theta^2 + \sum_{k=1}^N \ln x_k - \theta \sum_{k=1}^N x_k \end{aligned}$$

or

$$\frac{\partial L(\theta)}{\partial \theta} = \frac{2N\theta}{\theta^2} - \sum_{k=1}^N x_k = 0$$

and finally

$$\hat{\theta}_{ML} = \frac{2N}{\sum_{k=1}^N x_k}$$

Solution #4

2.23 From the theory, we have that the MAP estimate corresponds to the maximum of

$$\begin{aligned} Q(\mu) &\equiv \sum_{k=1}^N \ln p(x_k; \mu) + \ln p(\mu) \\ &= - \sum_{k=1}^N \frac{(x_k - \mu)^2}{2\sigma^2} + \ln \mu - \frac{\mu^2}{2\sigma_\mu^2} - \ln \sigma_\mu^2 \end{aligned}$$

or

$$\frac{\partial Q(\mu)}{\partial \mu} = - \sum_{k=1}^N \frac{\mu - x_k}{\sigma^2} + \frac{1}{\mu} - \frac{\mu}{\sigma_\mu^2} = 0$$

or

$$\left(\frac{N}{\sigma^2} + \frac{1}{\sigma_\mu^2}\right)\mu^2 - Z\mu - 1 = 0$$

and finally

$$\mu = \frac{Z}{2R} \left(1 + \sqrt{1 + \frac{4R}{Z^2}}\right)$$

since $\mu > 0$.

Solution #5

2.24

$$\begin{aligned} L(\theta) &= - \sum_{k=1}^N \frac{(\ln x_k - \theta)^2}{2\sigma^2} - \sum_{k=1}^N N \ln(\sigma\sqrt{2\pi}x_k) \\ \frac{\partial L(\mu)}{\partial \theta} = 0 &= - \sum_{k=1}^N \frac{\theta - \ln x_k}{\sigma^2} \Rightarrow \hat{\theta}_{ML} = \frac{1}{N} \sum_{k=1}^N \ln x_k \end{aligned}$$